

一类多重时滞系统的噪声统计 估值器和自适应滤波

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摘要

对于带有色观测噪声和多重滞后观测的线性离散系统, 本文提出了: 1) 次优递推滤波器; 2) 次优无偏递推极大后验 (MAP) 噪声统计估值器; 3) 自适应时变噪声统计估值器; 4) 自适应递推滤波器, 推广了 Sage 和 Husa [1] 和笔者 [2] 的结果。数值仿真例子证明了本文结果的有效性。本文结果可应用于地震信号去卷 [3]。

一、次优递推滤波器

考虑多重时滞线性离散系统

$$\mathbf{x}(k+1) = \sum_{i=0}^n A_i(k) \mathbf{x}(k-i) + \sum_{i=0}^m B_i(k) u(k-i) + w(k), \quad (1)$$

带有多重滞后的观测方程为

$$y(k) = \sum_{j=0}^M C_j(k) \mathbf{x}(k-j) + \alpha(k), \quad (2)$$

有色观测噪声 $\alpha(k)$ 为 N 阶自回归 (AR) 过程,

$$\alpha(k+1) = \sum_{j=0}^N D_j(k) \alpha(k-j) + v(k), \quad (3)$$

其中, \mathbf{x}, \mathbf{w} 是 p 维向量; y, α, v 是 r 维向量; u 是 s 维已知的输入向量; $A_i(k), B_i(k), C_i(k), D_i(k)$ 是已知的适当维矩阵; 而 $w(k)$ 与 $v(k)$ 是相互独立的高斯白噪声, 噪声统计分别为

$$E\mathbf{w}(k) = \mathbf{q}, \quad \text{var}\mathbf{w}(k) = Q; \quad E\mathbf{v}(k) = \mathbf{r}, \quad \text{var}\mathbf{v}(k) = R, \quad (4)$$

其中, E 是均值号, var 是方差号。

设基于观测 $(y(k), y(k-1), \dots, y(0))$ 对 $x(k)$ 的次优滤波估值为 $\hat{\mathbf{x}}(k)$, 相应的估值误差及其方差分别记为 $\tilde{\mathbf{x}}(k) = \mathbf{x}(k) - \hat{\mathbf{x}}(k)$, $P(k) = E(\tilde{\mathbf{x}}(k)\tilde{\mathbf{x}}^T(k))$, T 为转置号。为了导出系统 (1) — (4) 的次优递推滤波器, 本节暂时假定噪声统计是已知的, 且做如下的先验统计和初始条件的假定 [3]: 1) $\text{cov}(\mathbf{x}(k), w(j)) = 0, \text{cov}(\mathbf{x}(k), v(j)) = 0, k = 0, -1, \dots, -\theta, \forall j$;

2) $\text{cov}(\mathbf{x}(k), \mathbf{x}(j)) = 0, k \neq j, k, j = 0, -1, \dots, -\theta$; 3) 对 $k \geq \theta + 1$, 假定 $\tilde{\mathbf{x}}(k)$, $\tilde{\mathbf{x}}(k-1)$, \dots , $\tilde{\mathbf{x}}(k-\theta)$ 相互独立; 4) $\hat{\mathbf{x}}(i) = E\mathbf{x}(i)$, $\text{var}\mathbf{x}(i) = P(i)$, $i = 0, -1, \dots, -\theta$ 。其中 cov 是协方差号, $\theta = N + M + 1$ 。关于假定 1), 3) 的合理性解释和作用的讨论见文献[3]。

引入与 $y(k+1)$ 等价的新的观测过程 $z(k+1)$,

$$\mathbf{z}(k+1) = \mathbf{y}(k+1) - \sum_{i=0}^N D_i(k) \mathbf{y}(k-i), \quad k \geq \theta; \quad z(i) = y(i), \quad i = 0, 1, \dots, \theta. \quad (5)$$

由 (2) 和 (3) 式可得到带白噪声 $v(k)$ 的观测方程

$$\mathbf{z}(k+1) = \sum_{i=0}^{\theta} H_i(k+1) \mathbf{x}(k+1-i) + v(k), \quad (6)$$

其中, 矩阵 $H_i(k+1)$ 由合并 $\mathbf{x}(k+1-i)$ 的同类项得到

$$H_{i+1}(k+1) = C_{i+1}(k+1) - \sum_{j=0}^i D_{i-j}(k) C_j(k-i), \quad j = 0, 1, \dots, N+M, \quad i = 0, 1, \dots, N.$$

其中, 规定 $H_0(k+1) = C_0(k+1)$; $j > N$ 时, $D_j(k) = 0$; $j > M$ 时, $C_j(k-i) = 0$ 。

对于系统 (1) 和 (6) 式, 利用射影理论我们有如下次优递推滤波器^[4]:

$$\hat{\mathbf{x}}(k+1) = \hat{\mathbf{x}}(k+1/k) + K(k+1) \varepsilon(k+1), \quad (7)$$

$$\hat{\mathbf{x}}(k+1/k) = \sum_{i=0}^n A_i(k) \hat{\mathbf{x}}(k-i) + \sum_{i=0}^m B_i(k) u(k-i) + q, \quad (8)$$

$$\varepsilon(k+1) = \mathbf{y}(k+1) - \sum_{i=0}^n D_i(k) \mathbf{y}(k-i) - H_0(k+1) \hat{\mathbf{x}}(k+1/k) - \sum_{i=0}^{\theta-1} H_{i+1}(k+1) \hat{\mathbf{x}}(k-i) - r, \quad (9)$$

$$P(k+1/k) = \sum_{i=0}^n A_i(k) P(k-i) A_i^T(k) + Q, \quad (10)$$

$$\begin{aligned} K(k+1) = & \left[P(k+1/k) H_0^T(k+1) \right. \\ & + \sum_{i=0}^h A_i(k) P(k-i) H_{i+1}^T(k+1) \left. \right] \left[H_0(k+1) P(k+1/k) H_0^T(k+1) \right. \\ & + \sum_{i=0}^{\theta-1} H_{i+1}(k+1) P(k-i) H_{i+1}^T(k+1) + H_0(k+1) \left(\sum_{i=0}^h A_i(k) P(k-i) H_{i+1}^T(k+1) \right) \\ & + \left. \left(\sum_{i=0}^h A_i(k) P(k-i) H_{i+1}^T(k) \right)^T H_0^T(k+1) + R \right]^{-1}; \quad h = \min(n, \theta-1), \end{aligned} \quad (11)$$

$$\begin{aligned} P(k+1) = & \left[I - K(k+1) H_0^T(k+1) \right] P(k+1/k) \\ & - K(k+1) \sum_{i=0}^h H_{i+1}(k+1) P(k-i) A_i^T(k), \end{aligned} \quad (12)$$

其中, I 是单位阵, 初值为 $\hat{x}(i) = x_i$, $P(i) = P_i$, $i = 0, -1, \dots, -\theta$. $\varepsilon(k+1)$ 是 $z(k+1)$ 的次优新息过程, 带有统计特性

$$E\varepsilon(i) = 0, \quad \forall i \quad (13)$$

$$E[\varepsilon(i)\varepsilon^T(i)] = H_0(i)P(i/i-1)H_0^T(i) + \sum_{j=0}^{\theta-1} H_{j+1}(i)P(i-1-j)H_{j+1}^T(i)$$

$$\begin{aligned} & + H_0(i) \sum_{j=0}^h A_j(i-1) P(i-1-j) H_{j+1}^T(i) \\ & + \left[\sum_{j=0}^h A_j(i-1) P(i-1-j) H_{j+1}^T(i) \right]^T H_0^T(i) + R. \end{aligned} \quad (14)$$

上述结果包括普通Kalman滤波器^[1]和Tamura和Ueno^[3]的次优递推滤波器作为特殊情形。

二、次优无偏递推极大后验噪声统计估值器

对于系统(1)和(6)式, 当噪声统计未知时, 类似于Sage 和 Husa^[1]和作者^[2]的推导, 可得到基于信息($y(k+1), \dots, y(0)$)的次优无偏MAP噪声统计估值器

$$\hat{q}(k+1) = \frac{1}{k+1} \sum_{i=0}^k \left[\hat{x}(i+1) - \sum_{j=0}^h A_j(i) \hat{x}(i-j) - \sum_{j=0}^m B_j(i) u(i-j) \right], \quad (15)$$

$$\begin{aligned} \hat{Q}(k+1) = & \frac{1}{k+1} \sum_{i=0}^k \left[K(i+1) \varepsilon(i+1) \varepsilon^T(i+1) K^T(i+1) + P(i+1) \right. \\ & \left. - \sum_{j=0}^h A_j(i) P(i-j) A_j^T(i) \right], \end{aligned} \quad (16)$$

$$\begin{aligned} \hat{r}(k+1) = & \frac{1}{k+1} \sum_{i=0}^k \left[z(i+1) - H_0(i+1) \hat{x}(i+1/i) \right. \\ & \left. - \sum_{j=0}^{\theta-1} H_{j+1}(i+1) \hat{x}(i-j) \right], \end{aligned} \quad (17)$$

$$\begin{aligned} \hat{R}(k+1) = & \frac{1}{k+1} \sum_{i=0}^k \left[\varepsilon(i+1) \varepsilon^T(i+1) - H_0(i+1) P(i+1/i) H_0^T(i+1) \right. \\ & - \sum_{j=0}^{\theta-1} H_{j+1}(i+1) P(i-j) H_{j+1}^T(i+1) - H_0(i+1) \sum_{j=0}^h A_j(i) P(i-j) H_{j+1}^T(i+1) \\ & \left. - \left(\sum_{j=0}^h A_j(i) P(i-j) H_{j+1}^T(i+1) \right)^T H_0^T(i+1) \right]. \end{aligned} \quad (18)$$

事实上由(5),(7),(8),(9)和(13)式立刻有 $E\hat{q}(k+1) = q$, $E\hat{r}(k+1) = r$, 由(14)

式立刻有 $E\hat{R}(k+1) = R$ 。由(7), (8), (10), (11), (12)和(14)式易知 $E\hat{Q}(k+1) = Q$ 。这证明了无偏性。

易知上述估值器的递推形式为

$$\begin{aligned}\hat{q}(k+1) &= \frac{1}{k+1} \left[k\hat{q}(k) + \hat{x}(k+1) - \sum_{i=0}^n A_i(k)\hat{x}(k-i) - \sum_{i=0}^m B_i(k)u(k-i) \right], \\ (19) \end{aligned}$$

$$\begin{aligned}\hat{Q}(k+1) &= \frac{1}{k+1} \left[k\hat{Q}(k) + K(k+1)\varepsilon(k+1)\varepsilon^T(k+1)K^T(k+1) + P(k+1) \right. \\ &\quad \left. - \sum_{i=0}^n A_i(k)P(k-i)A_i^T(k) \right], \\ (20) \end{aligned}$$

$$\begin{aligned}\hat{r}(k+1) &= \frac{1}{k+1} \left[k\hat{r}(k) + \gamma(k+1) - \sum_{i=0}^N D_i(k)\gamma(k-i) \right. \\ &\quad \left. - \sum_{i=0}^{n-1} H_{i+1}(k+1)\hat{x}(k-i) - H_0(k+1)\hat{x}(k+1/k) \right], \\ (21) \end{aligned}$$

$$\begin{aligned}\hat{R}(k+1) &= \frac{1}{k+1} \left[k\hat{R}(k) + \varepsilon(k+1)\varepsilon^T(k+1) - H_0(k+1)P(k+1/k)H_0^T(k+1) \right. \\ &\quad \left. - \sum_{i=0}^{n-1} H_{i+1}(k+1)P(k-i)H_{i+1}^T(k+1) - H_0(k+1) \sum_{i=0}^h A_i(k)P(k-i)H_{i+1}^T(k+1) \right. \\ &\quad \left. - \left(\sum_{i=0}^h A_i(k)P(k-i)H_{i+1}^T(k+1) \right)^T H_0^T(k+1) \right]. \\ (22) \end{aligned}$$

初值取为 $\hat{q}(0) = q_0$, $\hat{r}(0) = r_0$, $\hat{Q}(0) = Q_0$, $\hat{R}(0) = R_0$ 。相应的自适应滤波器为在算法(7)–(12)式中把未知的噪声统计 q , Q , r , R 分别用由(19)–(22)式计算的在时刻 k 的估值 $\hat{q}(k)$, $\hat{Q}(k)$, $\hat{r}(k)$, $\hat{R}(k)$ 代替。

三、时变噪声统计自适应估值器

现在假定噪声均值和方差都是未知时变的: $Ew(k) = q(k)$, $\text{var}w(k) = Q(k)$, $Ev(k) = r(k)$, $\text{var}v(k) = R(k)$ 。
(23)

应用渐消记忆指数加权方法^[10]可得如下时变噪声统计自适应估值器:

$$\hat{q}(k+1) = (1-d_k)\hat{q}(k) + d_k \left[\hat{x}(k+1) - \sum_{i=0}^n A_i(k)\hat{x}(k-i) - \sum_{i=0}^m B_i(k)u(k-i) \right], \quad (24)$$

$$\begin{aligned}\hat{Q}(k+1) &= (1-d_k)\hat{Q}(k) + d_k \left[K(k+1)\varepsilon(k+1)\varepsilon^T(k+1)K^T(k+1) + P(k+1) \right. \\ &\quad \left. - \sum_{i=0}^n A_i(k)P(k-i)A_i^T(k) \right], \\ (25) \end{aligned}$$

$$\begin{aligned}\hat{r}(k+1) &= (1-d_k)\hat{r}(k) + d_k \left[y(k+1) - \sum_{i=0}^N D_i(k)y(k-i) - H_0(k+1)\hat{x}(k+1/k) \right. \\ &\quad \left. - \sum_{i=0}^{h-1} H_{i+1}(k+1)\hat{x}(k-i) \right],\end{aligned}\quad (26)$$

$$\begin{aligned}\hat{R}(k+1) &= (1-d_k)\hat{R}(k) + d_k \left[\varepsilon(k+1)\varepsilon^T(k+1) - H_0(k+1)P(k+1/k)H_0^T(k+1) \right. \\ &\quad \left. - \sum_{i=0}^{h-1} H_{i+1}(k+1)P(k-i)H_{i+1}^T(k+1) - H_0(k+1)\sum_{i=0}^h A_i(k)P(k-i)H_{i+1}^T(k+1) \right. \\ &\quad \left. - \left(\sum_{i=0}^h A_i(k)P(k-i)H_{i+1}^T(k+1) \right)^T H_0^T(k+1) \right],\end{aligned}\quad (27)$$

其中, $d_k = (1-b)/(1-b^{h+1})$, $0 < b < 1$, b 叫遗忘因子。

可证明在一定意义上上述自适应估值器对跟踪时变噪声统计而言是无偏的^[6]。

四、数值仿真例子

考虑如下单变量系统

$$x(k+1) = 0.7x(k) + 0.3x(k-1) + w(k), \quad (28)$$

$$y(k) = 0.8x(k) + 0.4x(k-1) + \alpha(k), \quad (29)$$

$$\alpha(k+1) = 0.8\alpha(k) + v(k), \quad (30)$$

其中, $w(k)$ 和 $v(k)$ 是零均值, 方差各为 0.02, 0.01 的高斯白噪声。在仿真过程中, 假定噪声统计是未知的, 并取初值为 $\hat{q}(0) = 0$, $\hat{Q}(0) = 0.001$, $\hat{r}(0) = 0$, $\hat{R}(0) = 0.004$; $\hat{x}(0) = 0.7$, $\hat{x}(-1) = 0.1$, $P(0) = P(-1) = 1$ 。应用本文算法仿真结果如图 1—3 所示, 说明了本文结果的有效性。

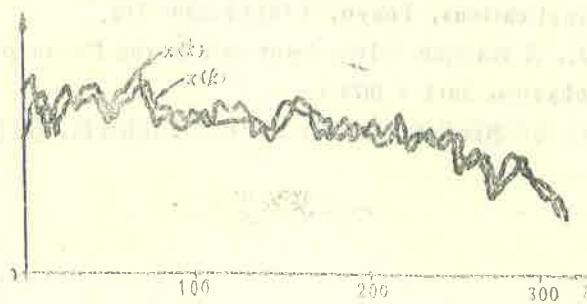
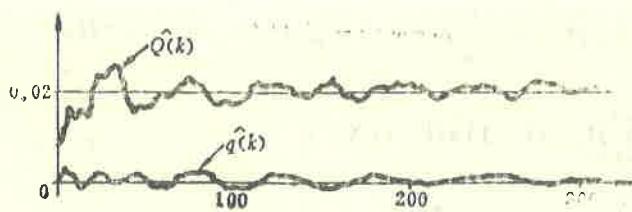
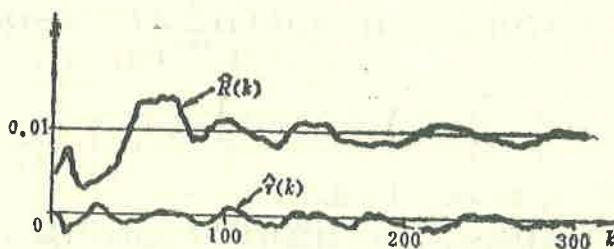


图 1 自适应滤波估值 $x(k)$ 与真实值 $x(k)$ 比较

图 2 估值 $\hat{q}(k)$, $\hat{Q}(k)$ 与真实值 q , Q 比较图 3 估值 $\hat{r}(k)$, $\hat{R}(k)$ 与真实值 r , R 比较

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Noise Statistics Estimators and Adaptive Filtering for a Class of Systems with Multiple Time Delays

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Abstract

For discrete linear systems with coloured observation noise and multiple delay observation, this paper presents (i) suboptimal recursive filter; (ii) suboptimal unbiased recursive maximum a-posteriori(MAP) estimators of unknown constant noise statistics; (iii) adaptive timevarying noise statistics estimators, and (iv) adaptive recursive filter. The results of [Sage and Husa^[1]] and author^[2] are extended. Numerical simulation example is given to show usefulness of proposed results. The results of this paper can be applied to the deconvolution of seismic signals^[5].